# Statistical analysis of experimental data Classification

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## Classification

- Event classification
- 2 Naive Bayes Classifier
  - 3 Fisher Linear Discriminant
  - 4 Nearest neighbors
- 5 Iterative procedure

#### 6 Homework



#### Problem



So far, we focused on the problem of extracting model parameters from the collected data sample. We used maximum likelihood approach (or  $\chi^2$  minimization, which is a special case).

However, what we often want to do is to "make choice", discriminate between two (or more) hypothesis based on the collected data.

We already addressed this problem (partially) when discussing limits (lecture 06) and consistency of the fit (lecture 07).

The general formulation of the problem: how to discriminate between two model hypothesis  $H_0$  and  $H_1$  based on the collected data D?

Common case:

 $H_0$  - Standard Model is valid,  $H_1$  - SM + additional BSM contribution

 $\boldsymbol{D}$  - the whole collected data sample, subset, or a single measurement



#### Neyman–Pearson Lemma

According to Neymann and Pearson, the optimal, "most powerful" method to discriminate between the two hypothesis is to look at likelihood ratio

$$Q(D) = \frac{L(D|H_1)}{L(D|H_0)}$$

When considering single measurements, making a cut on Q(x) is the optimal way to classify events. By using likelihood ratio, multi-dimensional measurements (whole events) are also presented as single number...

When we consider the whole sample of collected data, value of Q(D) is the best discriminant between the two hypothesis.

Still, one needs to compare the value of Q(D) resulting from the measurement, with the expected Q distributions for the two hypothesis.

## CL<sub>s</sub> method

The two hypothesis we consider in this case:

- H<sub>0</sub> Standard Model without Higgs contribution "background" only (b)
- $H_1$  SM with Higgs contribution "signal+background" (s+b) where we can consider different masses of the Higgs, m<sub>H</sub>

Instead of using Q, it is more convenient to use

 $q = -2\ln Q = -2\ell(D|H_1) + 2\ell(D|H_0) = \chi^2(D|H_1) - \chi^2(D|H_0)$ 

where:

- positive *q* values are expected for data more in agreement with background only hypothesis (*H*<sub>0</sub>)
- negative *q* values indicate that data are better described by signal+background hypothesis (*H*<sub>1</sub>)



## CL<sub>s</sub> method

Value of q from LEP,  $q_{dat}$ , was compared with distribution obtained with multiple Monte Carlo experiments for  $m_H = 115.6$  GeV.



We can define

$$CL_{s+b} = \int_{q_{dat}}^{+\infty} dq f^{H1}(q)$$

 $\Leftarrow$  indicated as blue area

$$CL_b = \int_{q_{dat}}^{+\infty} dq f^{H0}(q)$$

 $\Leftarrow$  indicated as red is  $1 - CL_b$ 



# - Fuw

## $\mathbf{CL}_{s}$ method

Experiments at LEP, running with energy up to  $\sqrt{s} = 210$  GeV, could only observe Higgs bosons with mass of up to about 118 GeV For higher masses, signal+background hypothesis ( $H_1$ ) becomes indistinguishable from background only one ( $H_0$ )...

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With tight event selection, experiments observed 4 candidate events with  $m_H^{rec} > 109 GeV$ . Expectations of background only hyposthesis: b = 1.2

In strictly frequentiest approach we could exclude (on 95%CL) not only the SM, but also all Higgs scenarios ( $H_1$ ) with  $m_H > 118 GeV!..$ 



# **CL**<sub>s</sub> method

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Frequentist approach gives us result which is correct (from statistical point of view) but not very useful... Too sensitive to background fluctuations?

Solution is to look for confidence level of  $H_1$  relative to  $H_0$ :

$$CL_s = \frac{CL_{s+b}}{CL_b}$$
  
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Probability of background hypothesis to result in  $N_{obs} \leq 7$  is 98.8%

 $\Rightarrow CL_{s} \text{ limit on number of signal events is 10.17} (95\% \text{ CL})$ almost the same as the Frequentist limit (CL<sub>s+b</sub>): 10.15

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Probability of background hypothesis to result in  $N_{obs} \leq 3$  is 64.7%

 $\Rightarrow CL_{s} \text{ limit on number of signal events is 5.40} \qquad (95\% \text{ CL}) \\ \text{only slightly higher than the Frequentist limit (CL_{s+b}): 4.75}$ 

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Probability of background hypothesis to result in  $N_{obs} \leq 1$  is 19.9%

 $\Rightarrow CL_{s} \text{ limit on number of signal events is 3.64} \qquad (95\% \text{ CL})$ significantly higher than the Frequentist limit (CL<sub>s+b</sub>): 1.74

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Probability of background hypothesis to result in  $N_{obs} = 0$  is 4.98%

 $\Rightarrow$  CL<sub>s</sub> limit on number of signal events is 3.00 (95% CL) while all signal hypothesis are excluded in Frequentist approach (CL<sub>s+b</sub>)!

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### $CL_s$ method





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#### **Problem definition**

The problem is similar to the one discussed in lecture 10: we want to discriminate between two model hypothesis  $H_0$  and  $H_1$  based on the collected data D.

Common case - interpretation of measurement results:

- H<sub>0</sub> Standard Model is valid,
- $H_1$  SM has to be extended by adding BSM contribution
- D the whole collected data sample

According to Neymann and Pearson, the optimal, "most powerful" method to discriminate between the two hypothesis is to look at likelihood ratio

$$Q(D) = \frac{L(D|H_1)}{L(D|H_0)}$$



#### **Problem definition**

The problem is similar to the one discussed in lecture 10: we want to discriminate between two model hypothesis  $H_0$  and  $H_1$  based on the collected data D.

Different case - classification of collected measurements:

- $H_0$  measurement can be attributed to the Standard Model,
- $H_1$  measurement is due to BSM contribution,
- *D* single measurement ("event" in HEP experiments)

According to Neymann and Pearson, the optimal, "most powerful" method to discriminate between the two hypothesis is to look at likelihood ratio

$$Q(D) = \frac{L(D|H_1)}{L(D|H_0)}$$

Fake gold coins have lower mass than the true ones.

We can select good coins (reject fake coins) by measuring the mass...



Fake coin contribution (5%) clearly visible (5 $\sigma$  separation)...



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Not so obvious when the mass difference is smaller...



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We can select good coins (reject fake coins) by measuring the mass...



Not so obvious when the mass difference is smaller... How to describe it?



1



#### Simple example

We use the Neymann-Pearson Lemma directly:

$$Q(m) = rac{L(m|H_1)}{L(m|H_0)}$$
 with  $rac{H_1 - \text{good coin}}{H_0 - \text{fake coin}}$ 

Assuming Gaussian uncertainties of the mass measurement

$$Q(m) = \frac{G(m; m_1, \sigma)}{G(m; m_0, \sigma)}$$



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$$Q(m) = \frac{G(m; m_1, \sigma)}{G(m; m_0, \sigma)}$$
$$= \exp\left[-\frac{1}{2}\left(\frac{m - m_1}{\sigma}\right)^2 + \frac{1}{2}\left(\frac{m - m_0}{\sigma}\right)^2\right]$$

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=  $\exp\left[\left(\frac{m_1-m_0}{\sigma^2}\right)\left(m-\frac{m_1+m_0}{2}\right)\right]$ 

Q is not very convenient to use, changes from very small to very large values. We already considered (LEP Higgs limits):

$$q(m) = -2 \log Q(m) = -2 \left( \frac{m_1 - m_0}{\sigma^2} \right) \left( m - \frac{m_1 + m_0}{2} \right)$$



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One can also consider discriminator function: -1 < y < +1

$$y(m) = \frac{1-Q}{1+Q} = \frac{L(m|H_1) - L(m|H_0)}{L(m|H_1) + L(m|H_0)}$$
  
=  $\tanh\left(-\frac{q(m)}{4}\right) = \tanh\left(\left(\frac{m_1 - m_0}{2\sigma^2}\right)\left(m - \frac{m_1 + m_0}{2}\right)\right)$ 



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Final selection should be based on a cut:  $y(m) > y_{cut}$  or  $Q(m) > Q_{cut}$  ...





Good vs fake coin discriminator function We expect  $y \rightarrow -1$  for fake coin,  $y \rightarrow +1$  for good coin





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Classification gets more and more difficult with decreasing  $\Delta m$ ...



#### **Classification errors**

O.Behnke et. al, Data Analysis in High Enegy Physics

Selecting the classification cut, two types of error need to be considered

	Reject <i>H</i> 0 (select as signal)	Accept <i>H</i> <sub>0</sub> (select as background)
<i>H</i> <sup>0</sup> is false (event is signal)	Right decision with probability $1 - \beta = power = efficiency$	Wrong decision; type II error with probability $\beta$
<i>H</i> <sup>0</sup> is true (event is background)	Wrong decision; type I error with probability $\alpha = size = significance$	Right decision with probability $1 - \alpha =$ background rejection

Probability of accepting fake		Probability of rejecting good	
$\alpha =$	$\int dm \ p(m H_0)$	$\beta =$	$\int dm \ p(m H_1)$
$y(m) > y_{cut}$		$y(m) < y_{cut}$	

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Discriminator function distribution

We expect  $y \to -1$  for fake coin,  $y \to +1$  for good coin



Large separation in  $\Delta m \Rightarrow$  very efficient classification possible



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#### Classification still possible, but error rate substantial



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Efficient classification "coin by coin" no longer possible...


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Good coin selection  $\Delta m = 4q \sigma = 1q$ -ake rejection 0.8  $1 - \alpha$ 0.6 0.4 0.2 0 0.2 0.4 0.6 0.8 0 Good efficiency  $1 - \beta$ 

In the realistic case, we can not have  $\alpha \rightarrow 0$  and  $\beta \rightarrow 0$  at the same time...



For both good and fake coins, efficiency depends on the assumed  $y_{cut}$  value. All possible choices on a Receiver-Operating-Characteristic curve:

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For both good and fake coins, efficiency depends on the assumed  $y_{cut}$  value. All possible choices on a Receiver-Operating-Characteristic curve:

Good coin selection  $\Delta m = 2q \sigma = 1q$ -ake rejection 0.8  $1 - \alpha$ 0.6 0.4 0.2 0 0.2 0.4 0.6 0.8 0 Good efficiency  $1 - \beta$ 

Optimal cut value strongly depends on the actual goal of the analysis...





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where we now consider classification of single measurement  $\mathbf{x}$ .

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However, this can be directly used only, if the likelihoods are known. For example (in particle physics), if we know the differential cross sections for the considered signal and background processes, and detector effects can be neglected...



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However, this can be directly used only, if the likelihoods are known. For example (in particle physics), if we know the differential cross sections for the considered signal and background processes, and detector effects can be neglected...

In most cases, we need to decide on the selection procedure based on the data (or pseud-data from Monte Carlo simulation) itself, try to use it to 'reconstruct' the likelihood ratio dependence on x...



#### **Bayes' Theorem**

#### refer to lecture 01

We can try to apply Bayes' Theorem to the classification problem. We can ask for the "probability" of the considered hypothesis H for given outcome x (data) of the measurement:

$$P(H|\mathbf{x}) = \frac{P(\mathbf{x}|H)}{P(\mathbf{x})} \cdot P(H)$$

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Let us assume that we know the probability density functions (properly normalized) for the two considered hypothesis:

$$p_0(\mathbf{x}) = P(\mathbf{x}|H_0) \quad p_1(\mathbf{x}) = P(\mathbf{x}|H_1)$$

and the expected fraction of events corresponding to  $H_1$ :  $f_1$ .



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and the expected fraction of events corresponding to  $H_1$ :  $f_1$ . Then:

$$P(H_1|\mathbf{x}) = \frac{f_1 \cdot p_1(\mathbf{x})}{f_1 \ p_1(\mathbf{x}) + (1 - f_1) \ p_0(\mathbf{x})}$$



#### refer to lecture 01



Single measurement (event) often corresponds to a set of observables:

 $\mathbf{x} = (x_1, x_2, \dots x_N)$ 

If N is large, it is difficult to reconstruct probability density function of  $\mathbf{x}$ .

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$$p_k^{(j)}(x_j) = P(x_j|H_k) = \int \cdots \int_{i \neq j} dx_i P(\mathbf{x}, H_k) \quad k = 1, 2$$



Single measurement (event) often corresponds to a set of observables:

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$$p_k^{(j)}(x_j) = P(x_j|H_k) = \int \cdots \int_{i \neq j} dx_i P(\mathbf{x}, H_k) \quad k = 1, 2$$

We can then apply the Bayes' Theorem to single variable distribution:

$$P(H_1|x_j) = \frac{f_1 \cdot p_1^{(j)}(x_j)}{f_1 \ p_1^{(j)}(x_j) + (1 - f_1) \ p_0^{(j)}(x_j)}$$





Assuming the absence of correlations between the observables, treating different observables as independent random variables, multi-deminsional pdf can be calculated as a product of variable pdfs.

Likelihood of hypothesis k for measured event **x** is then given by

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We can then construct the classifier based on the likelihood ratio:

$$\gamma(\mathbf{x}) = \frac{L_1(\mathbf{x})}{L_0(\mathbf{x}) + L_1(\mathbf{x})}$$

which should be equivalent to the Neyman-Pearson classifier. Assuming correlations can be neglected and in the limit of large training samples.

#### separation of 2D Gaussian distributions

We first use the train sample of events to reconstruct individual  $p_k^{(j)}$ . Training results can be then applied to the independent test sample...



Implementation of the Gaussian Naive Bayes Classifier in sklearn.

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#### Efficient classification can be obtained for uncorrelated variables.



Implementation of the Gaussian Naive Bayes Classifier in sklearn.



# Efficient classification can be obtained for uncorrelated variables. However, it is clearly far from optimal in case of correlations!



Implementation of the Gaussian Naive Bayes Classifier in sklearn.



# - **F.**..

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Classifier based on the linear combination of input variables:

$$F(\mathbf{x}; \mathbf{w}) = w_0 + \sum_{j=1}^{N} w_j x_j = w_0 + \mathbf{w} \cdot \mathbf{x}$$

Resulting decision boundaries,  $F(\mathbf{x}) = F_{cut}$ , are hyperplanes in N dim. How to find vector **w** giving best separation between two classes of events?



# (Behnke)

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Resulting decision boundaries,  $F(\mathbf{x}) = F_{cut}$ , are hyperplanes in N dim. How to find vector **w** giving best separation between two classes of events?

For the general case, to use numerical optimization procedure, one needs to define the "loss function". Possible choice: (similar to  $\chi^2$ )

$$L(\mathbf{w}) = \sum_{\text{events } i} \left[ t^{(i)} - y(F(\mathbf{x}^{(i)}; \mathbf{w})) \right]^2$$

where y is the decision ("activation") function (eg. step function or tanh),  $t^{(i)}$  is true class of training event  $\mathbf{x}^{(i)}$  (-1 for  $H_0$  and +1 for  $H_1$ ). Iterative procedure can be applied to minimize  $L(\mathbf{w})$ .



# (Behnke)

Weight vector **w** defines the direction, on which all events are projected. Projection "reduces" the N variable problem to single variable  $F(\mathbf{x})$ .





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Weight vector **w** defines the direction, on which all events are projected. Projection "reduces" the N variable problem to single variable  $F(\mathbf{x})$ .

If we assume Gaussian variable distributions, we can look at the direction which maximizes the relative distance between the two hypothesis in F:

$$D(\mathbf{w}) = \frac{(h_1 - h_0)^2}{\sigma_1^2 + \sigma_0^2}$$

 $h_k$  and  $\sigma_k^2$  are the expected values and variances of  $F(\mathbf{x})$  for hypothesis k:

$$h_k = \mathbb{E}(F(\mathbf{x})|H_k)$$
 and  $\sigma_k^2 = \mathbb{V}(F(\mathbf{x})|H_k)$ 



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How to relate  $h_k$  and  $\sigma_k^2$  to the properties of the **x** distribution for  $H_k$ :

$$\mu_k = \mathbb{E}(\mathbf{x}|H_k)$$
 and  $\mathbb{C}_{\mathbf{x}}$  ?

One can note that:

 $h_1 - h_0 = \mathbf{w} \cdot \boldsymbol{\mu}_1 - \mathbf{w} \cdot \boldsymbol{\mu}_0 = \mathbf{w} \cdot (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_0)$ 

and so the nominator of  $D(\mathbf{w})$  can be written as:

 $(h_1 - h_0)^2 = \mathbf{w}^{\mathsf{T}} \mathbb{B} \mathbf{w}$  where  $\mathbb{B} = (\mu_1 - \mu_0)(\mu_1 - \mu_0)^{\mathsf{T}}$ is the so-called between-class matrix.



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Introducing also the within-class matrix

$$\mathbb{W} = \mathbb{C}_{\mathbf{x}}^{(H_1)} + \mathbb{C}_{\mathbf{x}}^{(H_0)}$$

one can directly write F distance between two hypothesis in terms of w:

$$D(\mathbf{w}) = \frac{\mathbf{w}^{\mathsf{T}} \mathbb{B} \mathbf{w}}{\mathbf{w}^{\mathsf{T}} \mathbb{W} \mathbf{w}}$$





Matrices  $\mathbb{B}$  and  $\mathbb{W}$  depend only on **x** pdfs for the two hypothesis. They do not depend on the classifier weights **w**.

Maximum relative distance requirement corresponds to  $\nabla D(\mathbf{w}) = 0$  condition. The solution is:

$$\mathbf{w} = a \mathbb{W}^{-1} \left( \boldsymbol{\mu}_1 - \boldsymbol{\mu}_0 \right)$$

where a is an arbitrary scaling factor.  $w_0$  is the second free parameter

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$$\mathbf{w} = a \mathbb{W}^{-1} (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_0)$$

where *a* is an arbitrary scaling factor.  $w_0$  is the second free parameter When parameters of the pdfs are not known, they can be derived from the properties of the training sample:

$$\mathbf{w} = \mathbf{a} \left( \hat{\mathbb{C}}_{\mathbf{x}}^{(H_1)} + \hat{\mathbb{C}}_{\mathbf{x}}^{(H_o)} \right)^{-1} \left( \bar{\mathbf{x}}^{(H_1)} - \bar{\mathbf{x}}^{(H_0)} \right)$$

where  $\bar{\boldsymbol{x}}$  and  $\hat{\mathbb{C}}$  are the mean and covariance matrices for the training data



separation of 2D Gaussian distributions

Fisher discriminant takes variable correlations properly into account Different correlation coefficients  $\rho$  for well separated data



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#### separation of 2D Gaussian distributions

Fisher discriminant takes variable correlations properly into account Different relative separation for high correlation coefficient  $\rho = 0.9$ 





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# - **F**w

## Classification

- Event classification
- 2 Naive Bayes Classifier
- 3 Fisher Linear Discriminant
- 4 Nearest neighbors
- 5 Iterative procedure

#### 6 Homework

## This classifier refers directly to the Neymann and Pearson Lemma.

It is based on the expectation, that the likelihood ratio can be related to the ratio of the expected event densities:

$$Q(\mathbf{x}) = \frac{L(\mathbf{x}|H_1)}{L(\mathbf{x}|H_0)} = \frac{1}{N_1} \frac{dN_1}{d\mathbf{x}} \left(\frac{1}{N_0} \frac{dN_0}{d\mathbf{x}}\right)^{-1} = \frac{N_0}{N_1} \frac{dN_1}{dN_0}$$

where  $dN_k$  represent the expected number of events for hypothesis k, in a small variable space volume  $d\mathbf{x}$  (in the limit  $d\mathbf{x} \to 0$ ,  $N_k \to \infty$ )



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The idea is to replace the expected event densities  $dN_k$  by numbers of events in the actual data (training sample including  $H_0$  and  $H_1$  events):

$$dN_k(\mathbf{x}) \rightarrow n_k(\mathbf{x}) = \sum_{\mathbf{x}' \in \Delta(\mathbf{x})} \mathbf{x}' \subset H_k$$

The key point is how to define "neighborhood" region  $\Delta(\textbf{x})$  of point x



Two possible approaches are commonly used.

One can define  $\Delta(\mathbf{x})$  by specifying maximum distance d between points:

 $\Delta(\mathbf{x}) = \{x' : d(\mathbf{x}', \mathbf{x}) < R_{max}\}$ 

However,  $R_{max}$  has to be sufficiently large to always accept a sample of training events, also in the regions of lowest probability density... That is why this approach is not very efficient...



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This problem is "solved" in the "k nearest neighbors" (kNN) classification. We sort training events x' according to their distance from the test point x and take the closest k points:

$$\Delta(\mathbf{x}) = \{x' : d(\mathbf{x}', \mathbf{x}) < R(\mathbf{x})\} \text{ and } R(\mathbf{x}) : \sum_{\mathbf{x}'} 1 = k$$



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 and  $R(\mathbf{x}) : \sum_{\mathbf{x}'} 1 = k$ 

Where we still need to define the distance measure  $d(\mathbf{x}', \mathbf{x})$ ...





#### k Nearest Neighbors

We can then define our classifier as

$$\gamma(\mathbf{x}) = \frac{n_1(\mathbf{x})}{n_0(\mathbf{x}) + n_1(\mathbf{x})} = \frac{n_1(\mathbf{x})}{k}$$

where  $\gamma_{cut} = 0.5$  is assumed in most cases...



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For the distance measure, standard Euclidean metric is usually assumed:

$$d(\mathbf{x}',\mathbf{x}) = |\mathbf{x}'-\mathbf{x}| = \left(\sum_{j} (x'_{j}-x_{j})^{2}\right)^{\frac{1}{2}}$$

which, however, neglects differences in value scales of different variables and possible correlations between them



#### separation of 2D Gaussian distributions

Works already for k = 1: decision based on the nearest train event. Very large fluctuations, reflecting fluctuations in the training sample...



Implementation of the k Nearest Neighbors classifier in sklearn.

#### separation of 2D Gaussian distributions

#### Classification results more stable for larger k. Impact of fluctuations still visible...



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#### separation of 2D Gaussian distributions

Smooth classification boundary for k > 10, but details can be lost... Classification seems not to be optimal? (linear discriminator)



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separation of 2D Gaussian distributions

Precision of classification improves with size of training sample But it also gets more and more time consuming...



Implementation of the k Nearest Neighbors classifier in sklearn.

# Fw

#### kNN example

#### separation of 2D Gaussian distributions

Default procedure, assuming Euclidean metric, clearly fails when scales of the considered observables are very different...



Implementation of the k Nearest Neighbors classifier in sklearn.



## k Nearest Neighbors

To solve the problem of different variable scales, one could redefine considered set of variables to span the same numerical range or to have same variances. This however will still neglect possible correlations.



#### k Nearest Neighbors

To solve the problem of different variable scales, one could redefine considered set of variables to span the same numerical range or to have same variances. This however will still neglect possible correlations.

In the general case, distance measure properly reflecting properties of the data set should be used. Frequent choice:

$$d^{2}(\mathbf{x}',\mathbf{x}) = (\mathbf{x}'-\mathbf{x})^{\mathsf{T}} \mathbb{C}_{\mathbf{x}}^{-1} (\mathbf{x}'-\mathbf{x}) = \sum_{jk} (x_{j}'-x_{j}) (\mathbb{C}_{\mathbf{x}}^{-1})_{jk} (x_{k}'-x_{k})$$

where  $\mathbb{C}_{\boldsymbol{x}}$  is the covariance matrix of the measurement.

This is so-called "Mahalanobis distance" measure.

Similar to the calculation of the  $\chi^2$  value between two points

#### separation of 2D Gaussian distributions

Two-dimensional data set with large scale difference between variables With default distance measure (Euclidean metric)



Implementation of the k Nearest Neighbors classifier in sklearn.

separation of 2D Gaussian distributions

Two-dimensional data set with large scale difference between variables With Mahalanobis distance measure (including correlations)



Implementation of the k Nearest Neighbors classifier in sklearn.

separation of 2D Gaussian distributions

## Two-dimensional data set with large correlation between variables With default distance measure (Euclidean metric)



Implementation of the k Nearest Neighbors classifier in sklearn.



separation of 2D Gaussian distributions

Two-dimensional data set with large correlation between variables With Mahalanobis distance measure (including correlations)



Implementation of the k Nearest Neighbors classifier in sklearn.



separation of 2D Gaussian distributions

Two-dimensional data set with large correlation between variables With Mahalanobis distance measure (including correlations)



Results consistent with those obtained with Fisher linear discriminant.



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#### k Nearest Neighbors

While Naive Bayes and Fisher Linear Classifiers are based on modeling the likelihood distribution, nearest neighbors classifier is very general, can be used in (almost) any case. Example application:


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#### Homework

## Linear discriminant

In the standard approach, weights optimizing classifier based on the linear combination of input variables:

$$F(\mathbf{x};\mathbf{w}) = w_0 + \sum_{j=1}^N w_j x_j = w_0 + \mathbf{w} \cdot \mathbf{x}$$

are found based on the properties of the considered (training) samples.

However, the problem can be also solved without looking at the global properties, by minimizing the "loss function". Possible choice, "distance":

$$L(\mathbf{w}) = \sum_{\text{events } i} \left[ t^{(i)} - y(F(\mathbf{x}^{(i)}; \mathbf{w})) \right]^2$$

where y is the "activation function",  $t^{(i)}$  is true class of event  $\mathbf{x}^{(i)}$ . Many choices are possible for the activation function...





## **Activation function**



#### Source: Artificial Intelligence Wiki



Let activation function be defined in such a way that:

 $\lim_{x \to -\infty} y(x) = t(H_0) \quad \text{and} \quad \lim_{x \to +\infty} y(x) = t(H_1)$ 

so that for events far from division boundary (y = 0) we have:

$$y^{(i)} \equiv y\left(F(\mathbf{x}^{(i)};\mathbf{w})\right) \approx t^{(i)}$$

These events hardly contribute to the loss function.



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$$y^{(i)} \equiv y\left(F(\mathbf{x}^{(i)};\mathbf{w})\right) \approx t^{(i)}$$

These events hardly contribute to the loss function.

Events which are incorrectly classified contribute most to loss function. We can notice that the sign of  $y^{(i)} - t^{(i)}$  indicates the reason:

• if  $y^{(i)} - t^{(i)} > 0 \Rightarrow$  weights used were too large

• if  $y^{(i)} - t^{(i)} < 0 \Rightarrow$  weights used were too small



One can consider the iterative procedure of adjusting the weights:

$$\mathbf{w}^{(n+1)} = \mathbf{w}^{(n)} + \eta \sum_{i} \left( t^{(i)} - c^{(i)} \right) \cdot \mathbf{x}^{(i)}$$

where  $\eta$  is the learning rate parameter.

This approach was first proposed by M. Rosenblatt in 1958.

Weight correction can be applied on event by event basis (starting from the beginning when event loop completed) or calculating global correction for the whole sample.



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Surprisingly, with proper choice of  $\eta$  this procedure works, results in classification optimization, even without referring to the loss function...



N = 100

#### Perceptron Learning example

Example results for linear discriminant, starting from random weights:



Iterative procedure (dashed cyan) compared with Fisher discriminant (solid red)

A.F.Żarnecki



N = 1000

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Example results for linear discriminant, starting from random weights:



Iterative procedure (dashed cyan) compared with Fisher discriminant (solid red)

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## Classification

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#### Homework

#### Solutions to be uploaded by January 26.

You ordered 1000 silicon sensors,  $36 \times 24 \text{ mm}^2$ , in the factory. The order was processed on two machines (50% each) and it turns out that:

- all sensors from the first machine are working OK. They also have proper size tolerance  $\sigma_1 = 0.05$  mm
- the second machine was faulty, produced only faulty sensors and with worse size tolerance  $\sigma_2 = 0.25$  mm

Unfortunately, the two samples were mixed. When collecting your order, you can make selection based on the measured dimensions of the sensor and pay only for the sensors you select...

- find the optimal selection approach
- calculate the corresponding ROC curve
- what is the optimal selection cut, if you pay 10\$ for each collected sensor, and you can sell good ones for 12\$...

#### assume no correlation between the two dimensions

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